Exercises for Thursday Evening

1. Presented by: Flavia Cifarelli

Suppose that $y_t = x_t \beta + u_t$, with $u_t = 0.8u_{t-1} + \varepsilon_t$, and $x_t = \varepsilon_{t+1}$, where ε_t is iid(0,1). Let $\hat{\beta}$ denote the OLS estimator of β .

(a) Show that $\sqrt{T}(\hat{\beta} - \beta) \xrightarrow{d} N(0, V)$, and derive an expression for *V*.

(b) Noting that $(1-0.8L)u_t = \varepsilon_t$, the GLS estimator of β can be formed by regressing $(1-0.8L)y_t$ onto $(1-0.8L)x_t$. Derive the probability limit of the GLS estimator.

2. Presented by: Beatrice Retali

Suppose that $y_t = x_t \beta + u_t$, where $u_t = \phi u_{t-1} + \varepsilon_t$, $x_t = e_t + \theta e_{t-1}$, where ε_t and e_t are both i.i.d. with mean zero and variance σ_{ε}^2 and σ_{e}^2 , and ε_t and e_{τ} are independent for all *t* and τ . Let $\hat{\beta}$ denote the OLS estimator of β based on a sample of size *T*, and let \hat{u}_t denote the OLS residual.

(a) Show that $\sqrt{T}(\hat{\beta}-\beta) \xrightarrow{d} N(0,V) \sqrt{T}(\hat{\beta}-\beta) \xrightarrow{d} N(0,V)$ and derive an expression for V.

(b) Suppose that T = 100, $\hat{\beta} = 2.1$, $\frac{1}{100} \sum_{t=1}^{100} x_t^2 = 5$, $\frac{1}{99} \sum_{t=2}^{100} x_t x_{t-1} = 2.5$, $\frac{1}{98} \sum_{t=3}^{100} x_t x_{t-2} = 1.0$, $\frac{1}{100} \sum_{t=1}^{100} \hat{u}_t^2 = 4$, $\frac{1}{99} \sum_{t=2}^{100} \hat{u}_t \hat{u}_{t-1} = 3.6$, $\frac{1}{98} \sum_{t=3}^{100} \hat{u}_t \hat{u}_{t-2} = 3.1$, $\frac{1}{99} \sum_{t=2}^{100} x_t \hat{u}_{t-1} = 0.8$, $\frac{1}{99} \sum_{t=2}^{100} \hat{u}_t x_{t-1} = 0.2$. Use your result in (a) to construct an approximate 95% confidence interval for β .

3. Presented by: Laurant Ott

Suppose $y_t = \mu + u_t$ where u_t follows the stationary ARMA(1,1) process $u_t = \phi u_{t-1} + \varepsilon_t - \theta \varepsilon_{t-1}$ where $\varepsilon_t \sim \text{iid} (0, \sigma^2)$.

(a) Show that $\sqrt{T}(\overline{y} - \mu) \xrightarrow{a} N(0, V)$ and derive an expression for *V*. (Hint: use the formula for the ACGF for an ARMA(1,1) model to derive *V*.)

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(b) In a sample with T = 400, $\overline{y} = 5.4$. A researcher uses the data $y_t - \overline{y}$ to estimate the AR parameter and finds $\hat{\phi} = 0.66$, $\hat{\theta} = -0.81$ and $\hat{\sigma}_{\varepsilon} = 1.2$. Construct a 95% confidence interval for μ .

4. Presented by: Evert Reins

Suppose Y_t follows the AR(1) model $Y_t = \phi Y_{t-1} + \varepsilon_t$, where $\varepsilon_t \sim i.i.d., (0, \sigma^2)$. You are interested in forecasting Y_{T+2} given $Y_{1:T} = (Y_1, Y_2, ..., Y_T)$.

(a) Show that $E(Y_{T+2} | Y_{1:T}) = \gamma Y_T$, with $\gamma = \phi^2$.

(b) Let $\hat{\phi}$ denote the OLS estimator of ϕ , obtained from the regression of Y_{t+1} onto Y_t for t = 1, ..., T-1. Let $\hat{\gamma} = \hat{\phi}^2$ denote an estimator of γ . Show that $\sqrt{T}(\hat{\gamma} - \gamma) \xrightarrow{d} N(0, V_{\hat{\gamma}})$ and derive an expression for $V_{\hat{\gamma}}$.

(c) An alternative estimator of γ is obtained by regressing Y_{t+2} onto Y_t for t = 1, ..., T-2 using OLS. Call this estimator $\tilde{\gamma}$. Show that $\sqrt{T}(\tilde{\gamma} - \gamma) \xrightarrow{d} N(0, V_{\tilde{\gamma}})$ and derive an expression for $V_{\tilde{\gamma}}$.

(d) Is $\hat{\gamma}$ preferred to $\tilde{\gamma}$? Explain.

Some Additional Exercises

1. y_t follows the stationary AR(1) model $y_t = \phi y_{t-1} + \varepsilon_t$. A researcher wants to estimate ϕ . He cannot find data on y_t , but can find estimates of y_t based on survey. Let x_t denote the survey estimate of y_t , and suppose that $x_t = y_t + u_t$, where u_t is iid $(0, \sigma_u^2)$ and u_t is independent of ε_j for all *t* and *j*. The researcher estimates ϕ by regressing x_t onto x_{t-1} using x_{t-2} as an instrument.

- (a) Suppose that $\phi \neq 0$. Derive the asymptotic distribution of the IV estimator
- (b) Suppose $\phi = 0$, show that the IV estimator is not consistent. Derive the limiting distribution of the estimator.
- (c) How would you test $\phi = 0$?
- 2. Suppose that y_t follows the AR(1) process

$$y_t = \mu + u_t$$

$$u_t = \rho u_{t-1} + \varepsilon_t$$

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where ε_t is $iid(0,\sigma^2)$, and $|\rho| < 1$. Let

$$\overline{y} = T^{-1} \sum y_i$$

- (a) Show that $\overline{y} \xrightarrow{p} \mu$.
- (b) Show that $\sqrt{T}(\overline{y} \mu) \xrightarrow{d} N(0, V)$ and derive an expression for *V*.
- (c) Let $\hat{\lambda}_0 = \frac{1}{T} \sum_{i=1}^n (y_i \overline{y})^2$. Show that $\hat{\lambda}_0 \xrightarrow{p} \lambda_0$.

(d) Show that $\sqrt{T}(\hat{\lambda}_0 - \lambda_0) \xrightarrow{d} N(0, U)$, and derive an expression for U. (Feel free to make any additional assumptions necessary to show this result.)

3. Suppose that $y_t = x_t \beta + u_t$, where $u_t = \phi u_{t-1} + \varepsilon_t$, $x_t = e_t + \theta e_{t-1}$, where ε_t and e_t are both i.i.d. with mean zero and variance σ_{ε}^2 and σ_{e}^2 , and ε_t and e_{τ} are independent for all t and τ . Let $\hat{\beta}$ denote the OLS estimator of β based on a sample of size T, and let \hat{u}_t denote the OLS residual.

(a) Show that $\sqrt{T}(\hat{\beta} - \beta) \xrightarrow{d} N(0, V)$ and derive an expression for V.

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(b) Suppose that
$$T = 100$$
, $\hat{\beta} = 2.1$, $\frac{1}{100} \sum_{t=1}^{100} x_t^2 = 5$, $\frac{1}{99} \sum_{t=2}^{100} x_t x_{t-1} = 2.5$, $\frac{1}{98} \sum_{t=3}^{100} x_t x_{t-2} = 1.0$,
 $\frac{1}{100} \sum_{t=1}^{100} \hat{u}_t^2 = 4$, $\frac{1}{99} \sum_{t=2}^{100} \hat{u}_t \hat{u}_{t-1} = 3.6$, $\frac{1}{98} \sum_{t=3}^{100} \hat{u}_t \hat{u}_{t-2} = 3.1$, $\frac{1}{99} \sum_{t=2}^{100} x_t \hat{u}_{t-1} = 0.8$,
 $\frac{1}{99} \sum_{t=2}^{100} \hat{u}_t x_{t-1} = 0.2$. Construct a 95% confidence interval for β .

(c) An alternative to OLS in this problem is GLS. Explain how you would construct the GLS estimator. Is GLS preferred to OLS in this situation? Explain.

4. Suppose that a sequence of random variables y_t is generated by the model:

$$y_t = \mu + u_t$$
$$u_t = \varepsilon_t + \theta \varepsilon_{t-1}$$

where $\varepsilon_t \sim Niid(0, \sigma^2)$. From a realization of size T = 100, I calculate

$$\overline{y} = \frac{1}{T} \sum_{t=1}^{T} y_t = 4$$
$$\sum_{t=2}^{T} (y_t - \overline{y})(y_{t-1} - \overline{y}) = 30$$
$$\sum_{t=1}^{T} (y_t - \overline{y})^2 = 100$$

- (a) Show that \overline{y} is the ordinary least squares estimate of μ .
- (b) Calculate an estimate of the variance of the OLS estimator of μ .
- (c) Construct a 95% confidence interval for μ .
- 5. Suppose that y_t follows the AR(1) process

$$y_t = \mu + u_t$$
$$u_t = \rho u_{t-1} + \varepsilon_t$$

where ε_t is $iid(0,\sigma^2)$, and $|\rho| < 1$. Let

$$\hat{\mu}^{ols} = T^{-1} \sum y_t$$

(e) Show that $\sqrt{T}(\hat{\mu}^{ols} - \mu) \xrightarrow{d} N(0, V_{ols})$ and derive an expression for V_{ols} .

(b) Suppose that $u_0 = 0$. Explain how you would construct the GLS estimator of μ .

(c) Let $\hat{\mu}^{gls}$ denote the GLS from part c. Show that $\sqrt{T}(\hat{\mu}^{gls} - \mu) \xrightarrow{d} N(0, V_{gls})$ and derive an expression for V_{gls} .

(d) Is the GLS estimator better that the OLS estimator? Explain.

(e) Suppose that T = 100, and $\hat{\mu}^{ols} = 2$. Let $\hat{u}_t = y_t - \hat{\mu}^{ols}$. The regression of \hat{u}_t onto \hat{u}_{t-1} yields a regression coefficient of 0.4 and the standard error of the regression is 1.1. Construct a 95% confidence interval for μ .

6. Suppose that y and x follow the process

$$y_t = ax_{t-1} + \varepsilon_t$$
$$x_t = \phi x_{t-1} + v_t$$

where $v_t = \varepsilon_t + e_t$, where ε and e are mutually independent iid (0,1) processes. Suppose that $|\phi| < 1$. Let $\hat{\alpha}$ and $\hat{\phi}$ denote the OLS estimators of α and ϕ .

(a) Derive asymptotic distribution of the 2×1 vector $(\hat{\alpha}, \hat{\phi})$.

(b) In a sample of 100 observations $\hat{\alpha} = 1.3$ and $\hat{\phi} = 0.72$. Derive a 95% confidence interval for α .

7. Suppose that Y_t follows a stationary AR(1), and you have data Y_t , t = 1, ..., 100, and find $\overline{Y} = T^{-1} \sum_{t=1}^{T} Y_t = 18.1$, $\hat{\lambda}_0 = T^{-1} \sum_{t=1}^{T} (Y_t - \overline{Y})^2 = 2.4$, and $\hat{\lambda}_1 = (T-1)^{-1} \sum_{t=1}^{T} (Y_t - \overline{Y})(Y_{t-1} - \overline{Y}) = 1.0$. Construct a 95% confidence interval for $\mu = E(Y)$. (Hint: Proceed in three steps. Step 1: Think about the appropriate asymptotic covariance matrix for $\sqrt{T}(\overline{Y} - \mu)$. Step 2: Think about how you would estimate the covariance matrix from the data given in the problem. Step 3: Construct the confidence interval.)

- 9. Suppose that y_t is iid $(0,\sigma^2)$ with $E(y_t^4) = \kappa$. Let $\hat{\sigma}^2 = \frac{1}{T} \sum_{t=1}^T y_t^2$.
 - (a) Show that $\sqrt{T}(\hat{\sigma}^2 \sigma^2) \xrightarrow{d} N(0, V_{\hat{\sigma}^2})$ and derive an expression for $V_{\hat{\sigma}^2}$.

(b) Suppose that T = 100, $\hat{\sigma}^2 = 2.1$ and $\frac{1}{T} \sum_{t=1}^{T} y_t^4 = 16$. Construct a 95% confidence interval for σ , where σ is the standard deviation of y.

10. Suppose that y_t follows the AR(1) process $y_t = \phi y_{t-1} + \varepsilon_t$, where $\varepsilon_t \sim \text{iidN}(0, \sigma_{\varepsilon}^2)$. Let $\hat{\phi}$ denote the OLS estimator, $\hat{\varepsilon}_t = y_t - \hat{\phi} y_{t-1}$ denote the OLS residual, and $\hat{\sigma}_{\varepsilon}^2 = \frac{1}{T} \sum_{t=1}^T \hat{\varepsilon}_t^2$ denote the estimator of σ_{ε}^2 .

(a) Suppose that $|\phi| < 1$ and $\{y_t\}$ is stationary. Show that $\hat{\sigma}_{\varepsilon}^2 \xrightarrow{p} \sigma_{\varepsilon}^2$

(b) Suppose that $\phi = 1$ and $y_0 = 0$. Show that $\hat{\sigma}_{\varepsilon}^2 \xrightarrow{p} \sigma_{\varepsilon}^2$. (Hint: Answer this after studying the unit root autoregression.)

11. (a) Suppose that y_t , for t = 1, ..., T is generated by the model: $y_t = \mu + u_t$, where $u_t = \varepsilon_t - \theta \varepsilon_{t-1}$, where $\varepsilon_t \sim i.i.d$. with mean zero and variance σ^2 .

(a.1) Show that $\sqrt{T}(\overline{Y} - \mu) \xrightarrow{d} N(0, V)$ and derive an expression for *V*.

(a.2) Propose an estimator for V and sketch a proof showing that the estimator is consistent.

(b) Now suppose that y_t follows the same model as in (a) but the parameter μ undergoes a change during the sample. That is for $1 \le t \le \tau$, the parameter takes on the value μ_1 and for $\tau+1 \le t \le T$ the parameter takes on the value μ_2 . (The parameters θ and σ^2 do not change.) Let $\overline{Y}_1 = \frac{1}{\tau} \sum_{t=1}^{r} y_t$ and $\overline{Y}_2 = \frac{1}{T-\tau} \sum_{t=\tau+1}^{T} y_t$. Let $\rho = \tau/T$. (b.1) Suppose that ρ is fixed as $T \to \infty$. Show that $\sqrt{T} \left[\left(\overline{Y}_1 - \overline{Y}_2 \right) - \left(\mu_1 - \mu_2 \right) \right] \xrightarrow{d} N(0, \Omega)$ and derive an expression for Ω .

(b.2) Discuss how you would construct an approximate 95% confidence interval for $\mu_1 - \mu_2$ using the data y_t , t = 1, ..., T.

(b.3) Describe how you would test the null hypothesis that $\mu_1 = \mu_2$ if you know the break date τ .

12. Suppose that y_t follows a stationary Gaussian zero-mean ARMA(1,1) process, $(1-\rho L)y_t = (1-\theta L)\varepsilon_t$. Let $\hat{\rho}$ denote the IV estimator of ρ obtained by "regressing" y_t onto y_{t-1} using y_{t-2} as an instrument. Let *T* denote the sample size.

- (a) Suppose that the true value of ρ is $\rho = 0.5$ and the true value of θ is $\theta = 0.1$. (i) Show that $\hat{\rho}$ is consistent.
 - (ii) Derive the asymptotic distribution of $\hat{\rho}$.
- (b) Suppose that the true value of ρ is $\rho = 0.5$ and the true value of θ is $\theta = 0.5$.
 - (i) Is $\hat{\rho}$ is consistent? Explain.
 - (ii) Derive the asymptotic distribution of $\hat{\rho}$.

13. Consider the IV model with L = K = 1, and conditionally homoskedastic errors. (That is $y_t = z_t \delta + \varepsilon_t$, x_t is a instrument, and $var(\varepsilon_t x_t) = \sigma_{\varepsilon}^2 \sigma_{\chi}^2$.)

In a sample of size 200, the following sample moment matrices are calculated:

 $S_{yy} = 4.33$ $S_{zy} = 1.87$ $S_{zz} = 1.12$ $S_{xy} = 0.56$ $S_{xz} = 0.58$ $S_{xx} = 1.06$

(a) Compute $\hat{\delta}$ and its standard error (based on the usual large sample approximation).

(b) Compute a 95% confidence for δ using the usual large sample approximation to the distribution of $\hat{\delta}$.

(c) Compute the *F*-statistic from the regression of z onto x. What is the *p*-value for this statistic? Is there evidence that x is a weak instrument?

(d) Use the Anderson-Rubin method to compute a 95% confidence interval for δ . How does this confidence interval compare to the confidence interval that you computed in (a). Is this what you have expected given the *F*-statistic in (c)? Explain.

14. y_t follows the stationary AR(1) model $y_t = \phi y_{t-1} + \varepsilon_t$. A researcher wants to estimate ϕ . He cannot find data on y_t , but can find estimates of y_t based on survey. Let x_t denote the survey estimates of y_t , and suppose that $x_t = y_t + u_t$, where u_t is iid(0, σ_u^2) and u_t is independent of ε_j for all *t* and *j*. The researcher estimates ϕ by regressing x_t onto x_{t-1} using x_{t-2} as an instrument.

(a) Suppose that $\phi \neq 0$. Derive the asymptotic distribution of the IV estimator (b) Suppose $\phi = 0$, show that the IV estimator is not consistent. Derive limiting distribution.

(c) How would you test $\phi = 0$?

15. Suppose that y_t follows the MA(1) process $y_t = (1 - \theta L)\varepsilon_t$, where $\varepsilon_t \sim iid(0, \sigma_{\varepsilon}^2)$ and has as many higher-order moments as needed to answer the questions below.

- (a) Show that the variance of y is given by $\sigma_y^2 = \sigma_\varepsilon^2 (1 + \theta^2)$.
- (b) Let $\hat{\sigma}_y^2 = T^{-1} \sum_{t=1}^{T} y_t^2$. (b.i) Show that $\hat{\sigma}_y^2 \xrightarrow{p} \sigma_y^2$ (b.ii) Show that $\sqrt{T} (\hat{\sigma}_y^2 - \sigma_y^2) \xrightarrow{d} N(0, V)$ and derive an expression for *V* in terms of moments of *y*.
- (c) Using a sample of size T = 100, I find $\hat{\sigma}_y^2 = 1.52$, $T^{-1} \sum_{t=1}^T (y_t^2 \hat{\sigma}_y^2)^2 = 4.8$, $T^{-1} \sum_{t=1}^T (y_t^2 - \hat{\sigma}_y^2) (y_{t-1}^2 - \hat{\sigma}_y^2) = 1.1$, $T^{-1} \sum_{t=1}^T (y_t^2 - \hat{\sigma}_y^2) (y_{t-2}^2 - \hat{\sigma}_y^2) = -0.4$, $T^{-1} \sum_{t=1}^T y_t y_{t-1}^3 = 3.6$, and $T^{-1} \sum_{t=1}^T y_t^3 y_{t-1} = 3.1$. Construct a 95% confidence interval for σ_y^2 .

16.. y_t follows the stationary AR(1) model $y_t = \phi y_{t-1} + \varepsilon_t$. A researcher wants to estimate ϕ . He cannot find data on y_t , but can find estimates of y_t based on survey. Let x_t denote the survey estimates of y_t , and suppose that $x_t = y_t + u_t$, where u_t is iid(0, σ_u^2) and u_t is independent of ε_j for all *t* and *j*. The researcher estimates ϕ by regressing x_t onto x_{t-1} using x_{t-2} as an instrument.

- (a) Suppose that $\phi \neq 0$. Derive the asymptotic distribution of the IV estimator
- (b) Suppose $\phi = 0$, show that the IV estimator is not consistent. Derive limiting distribution.
- (c) How would you test $\phi = 0$?