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# Credibility of Confidence Sets in Nonstandard Econometric Problems

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April 2016

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## Motivation

- Confidence sets are most common way of describing parameter uncertainty in econometrics
- Definition of confidence sets does not ensure that they do so in a "reasonable" fashion:
  - Observe  $X \sim \mathcal{N}(\theta, 1)$ ,  $\theta > 0$
  - Use  $CS = [X - 1.96, X + 1.96] \cap (0, \infty)$  (inversion of UMPU tests)
  - Empty with positive probability
  - "Too short" if  $X \approx -1.95$

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## Obviously "Unreasonable" Non-Empty CS

- Observe  $X = (Y, S)$ , where

$$P(S = 1) = P(S = 5) = 1/2$$
$$Y|S \sim \mathcal{N}(\theta, S^2)$$

- Standard 95%  $CS_0$  is  $[Y \pm 1.96S]$ . But

$$CS(X) = \begin{cases} [Y \pm 2.58S] & \text{if } S = 1 \\ [Y \pm 1.70S] & \text{if } S = 5 \end{cases}$$

has shorter expected length

- CS is "unreasonably" short whenever  $S = 5$  (failure to condition on the ancillary statistic  $S$ ?)

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## Formalization of "Reasonable" CS

- After realization  $X = x$ , inspector can choose to bet against the event that  $\theta \in \text{CS}(x)$ , winning 1 if right, and losing  $\alpha/(1 - \alpha)$  if wrong
- Can inspector generate positive average winnings from repeated draws, for all possible values of true parameter?
  - Surely yes if CS is empty with positive probability
  - Also if CS overstates the accuracy, as in conditional normal example
- Central idea of this paper: Use this betting scheme to study reasonableness of confidence sets in non-standard problems of recent interest
  1. Magnitude of maximal expected winnings is quantitative measure of "unreasonableness" of previously suggested CS
  2. Construction of attractive "reasonable" (=bet-proof) CS

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## Literature

- Previous results on betting and conditioning
  - Relevant/recognizable subsets and betting: Fisher (1956), Buehler (1959), Wallace (1959), Cornfield (1969), Pierce (1973), Robinson (1977), ...
  - Condition on ancillary statistics: Fisher (1956), Basu (1959), Basu (1964), ...
- Interesting to revisit these ideas, as issues are potentially important in nonstandard problems of applied interest
- We make (analytical and numerical) progress in construction of bet-proof CS

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## Remainder of Talk

1. Quantification of unreasonableness of popular CS in nonstandard problems
  - (a) AR(1) coefficient near unity; (b) weak instruments
2. Characterization of bet-proof sets
3. Construction of attractive bet-proof sets: Theory
4. Construction of attractive bet-proof sets: Applications
  - (a) AR(1) coefficient near unity; (b) weak instruments
5. Conclusion

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## Quantifying Unreasonableness

- For any weighting function  $\pi$  on the parameter space, one can characterize the betting strategy that maximizes weighted average expected winnings, subject to non-negative expected winnings for each  $\theta$
- Also investigate magnitude of maximal expected winnings under less favorable payoffs: inspector still wins 1 if right, but now loses  $\alpha'/(1 - \alpha')$  for wrong objections, where  $\alpha' > 0.05$
- Upper bound on expected winnings for all problems and  $1 - \alpha$  CS is  $\alpha = 0.05$

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## AR(1) Coefficient Near Unity

- Consider AR(1)

$$y_t = \mu + u_t$$

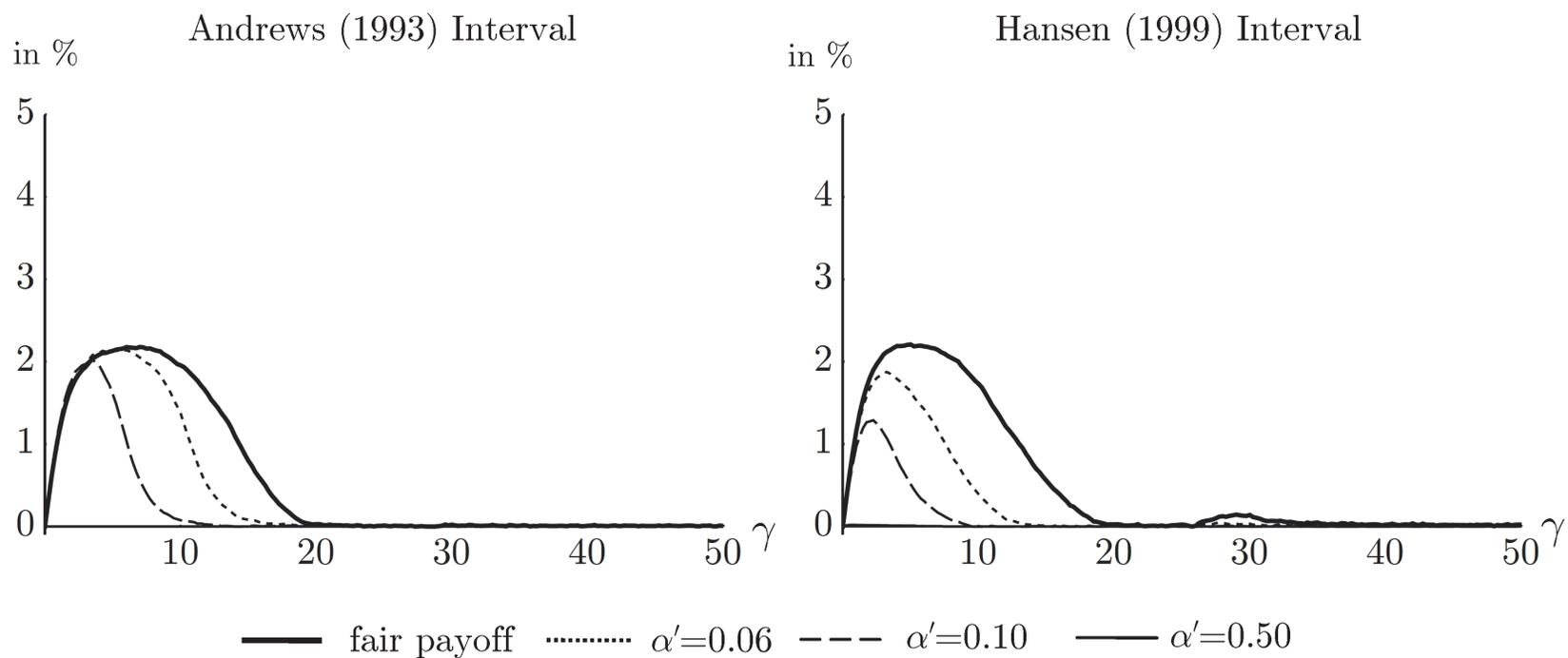
$$u_t = \rho u_{t-1} + \varepsilon_t$$

where  $\varepsilon_t \sim iid\mathcal{N}(0, \sigma^2)$  and  $u_0 \sim \mathcal{N}(0, \sigma^2/(1 - \rho^2))$

- Analyze inference under local-to-unity asymptotics  $\rho = \rho_T = 1 - \gamma/T$
- Two 95% CS for  $\gamma$  by inverting equal-tailed tests based on
  - OLS estimator  $\hat{\rho}$  (Andrews (1993))
  - OLS t-statistic based on  $\hat{\rho}$  (Hansen (1999))
- Specify  $\pi(\gamma) \propto (100 + \gamma)^{-1.1} \mathbf{1}[\gamma \geq 0]$

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# AR(1) Coefficient Near Unity



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## Weak Instruments

- In Chamberlain's (2005) reparamterization, weak instrument inference with one endogenous variable and one instrument is identical to inference about  $\tan(\phi)$  when observing

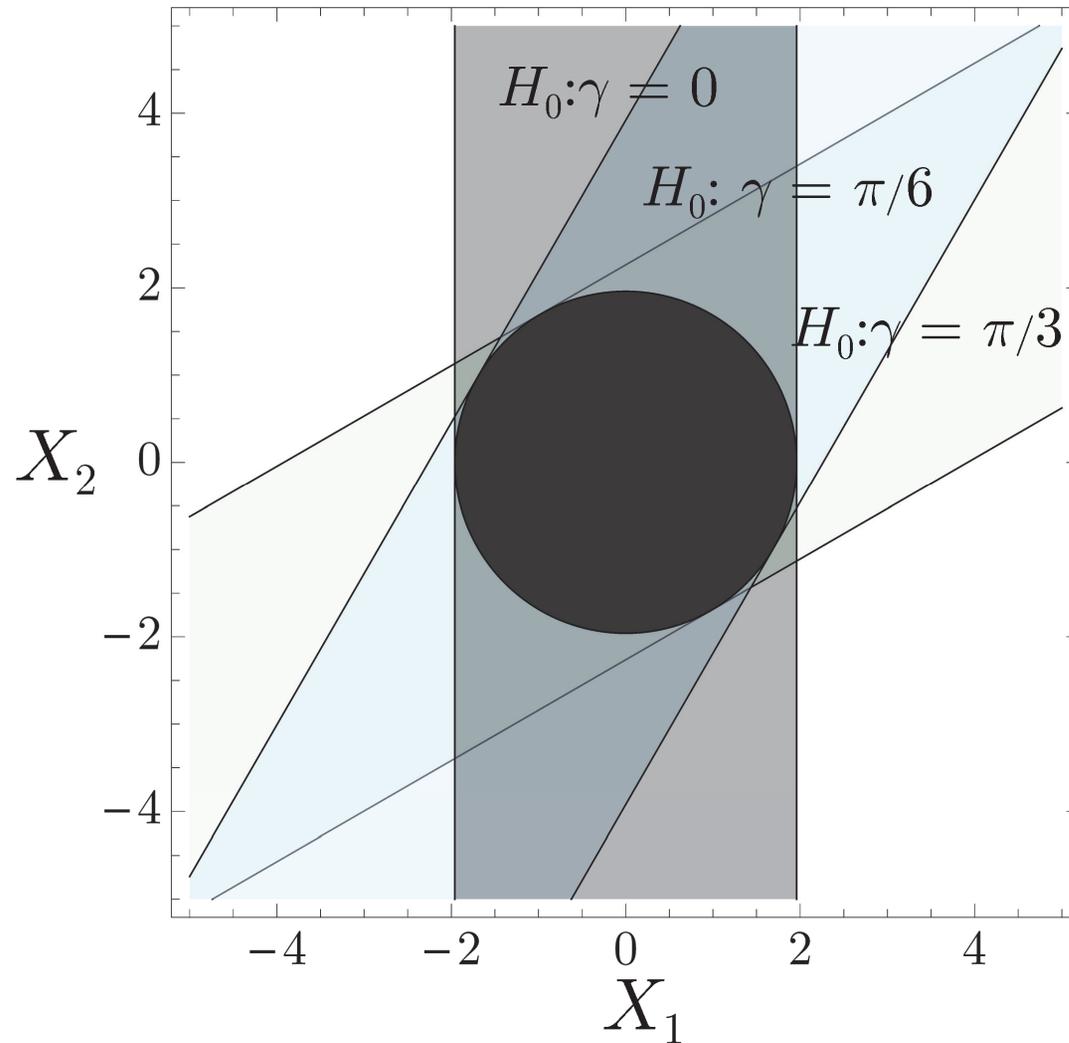
$$X = \begin{pmatrix} X_1 \\ X_2 \end{pmatrix} \sim \mathcal{N} \left( \begin{pmatrix} \rho \sin \phi \\ \rho \cos \phi \end{pmatrix}, I_2 \right)$$

where  $\phi \in [0, 2\pi]$  and  $\rho \geq 0$ .

- $\tan(\phi)$  is a one-to-one function of  $\gamma = (\phi \bmod \pi) \in [0, \pi]$
- 5% level Anderson-Rubin test of  $H_0 : \gamma = \gamma_0$  rejects if  $|X_1 \cos \gamma_0 - X_2 \sin \gamma_0| > 1.96$ 
  - $\Rightarrow$  Inverting the AR test yields a similar 95% CS
  - $\Rightarrow$  If  $\|X\| < 1.96$ , then  $\text{AR-CS}(X) = [0, \pi]$  (no value is excluded)

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# Acceptance Regions



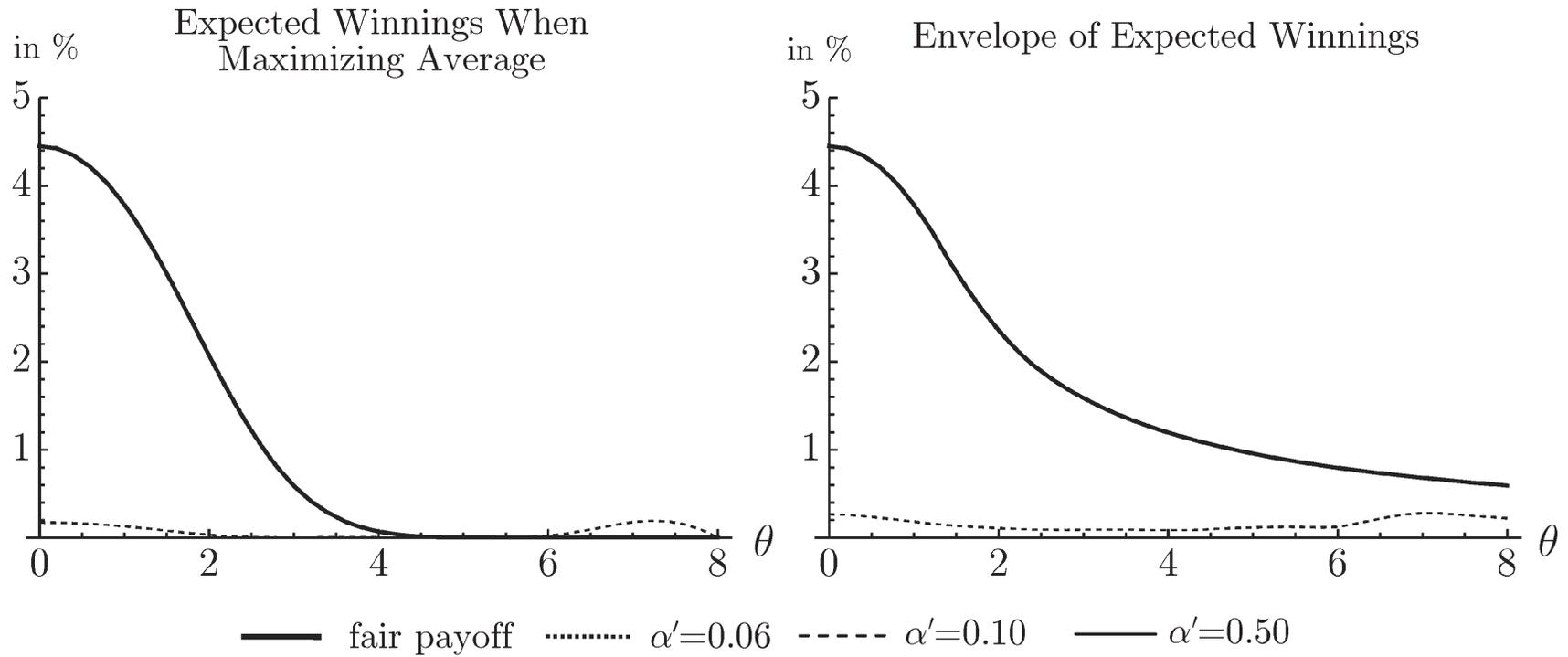
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## General Result

- If a *similar* CS is equal to "all parameter values" with positive probability, then betting whenever " $CS \neq \text{all values}$ " yields uniformly positive expected winnings
  - Intuition:
    - Probability of coverage conditional on " $CS = \text{all values}$ " is unity
    - Similarity forces that average probability of coverage is equal to  $1 - \alpha$
- ⇒ Similar CS must undercover conditional on " $CS \neq \text{all values}$ "

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# Weak Instruments



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## Characterization of Bet-Proof CS

- Previously known Theorem: Under weak technical conditions, a level  $(1 - \alpha)$  CS is bet-proof (there is no betting strategy with positive average winnings uniformly over the parameter space) if and only if CS is a superset of a  $(1 - \alpha)$  credible set for some prior.

⇒ Asymptotic equivalence of frequentist CS and Bayesian HPDs in standard parametric models implies that frequentist CS are nearly bet-proof

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## Construction of Bet-Proof Sets

- By previous result: any bet-proof set must be a superset of a credible set relative to some prior
- One approach: Can one find a prior that turns a specific level  $1 - \alpha$  credible set (such as, say, HPD) into a confidence set?
  - Yes, see Müller and Norets (2015, forthcoming JASA)
  - But might be conservative in presence of nuisance parameters
- Alternative pursued here: Guarantee bet-proofness by always including (some specific) credible set relative to exogenous prior

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## Construction of Bet-Proof Sets, Ctd

- Given the exogenous prior  $\pi^0$ , construct credible set  $S_0(x)$  from data  $x$
- Enlarge  $S_0$  to ensure coverage by a minimal amount in a weighted average length sense

$$\begin{aligned} \min_S \int E_\theta[\text{vol}(S(X))] d\pi^0(\theta) \\ \text{s.t. } P_\theta(\theta \in S(X)) \geq 1 - \alpha \\ \text{and } S(x) \supset S_0(x) \text{ for all } x \end{aligned}$$

$\Rightarrow S$  inverts weighted average power maximizing tests (Pratt 1961), subject to not rejecting if  $\theta_0 \in S_0(x)$

- Paper contains details on extension to invariant problems, and problems with nuisance parameters

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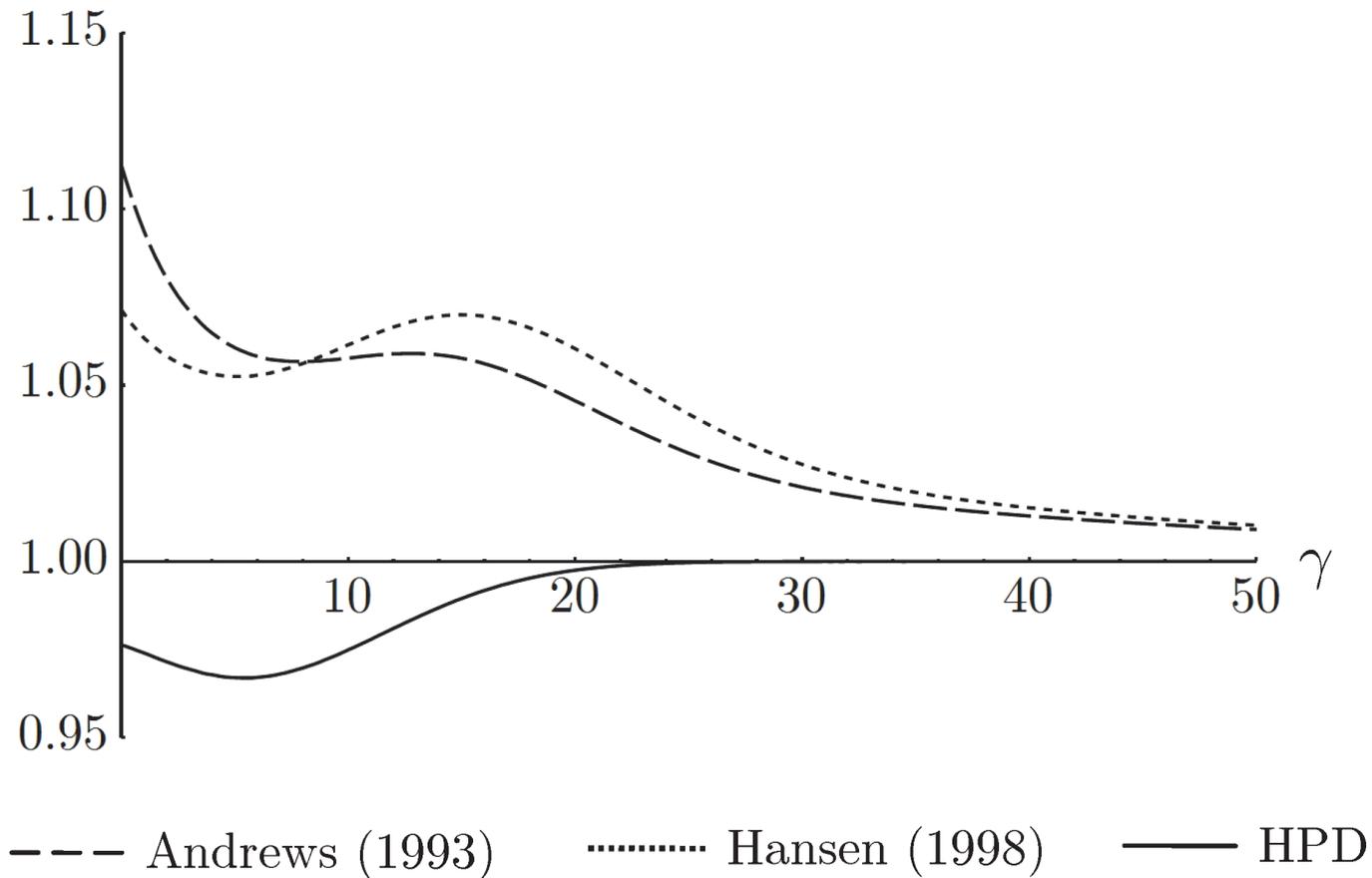
## AR(1) Coefficient Near Unity

- Impose translation invariance
- Use same weighting function/prior as before:

$$\pi^0(\gamma) \propto (100 + \gamma)^{-1.1} \mathbf{1}[\gamma \geq 0]$$

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## AR(1) Expected Lengths Relative to $S$



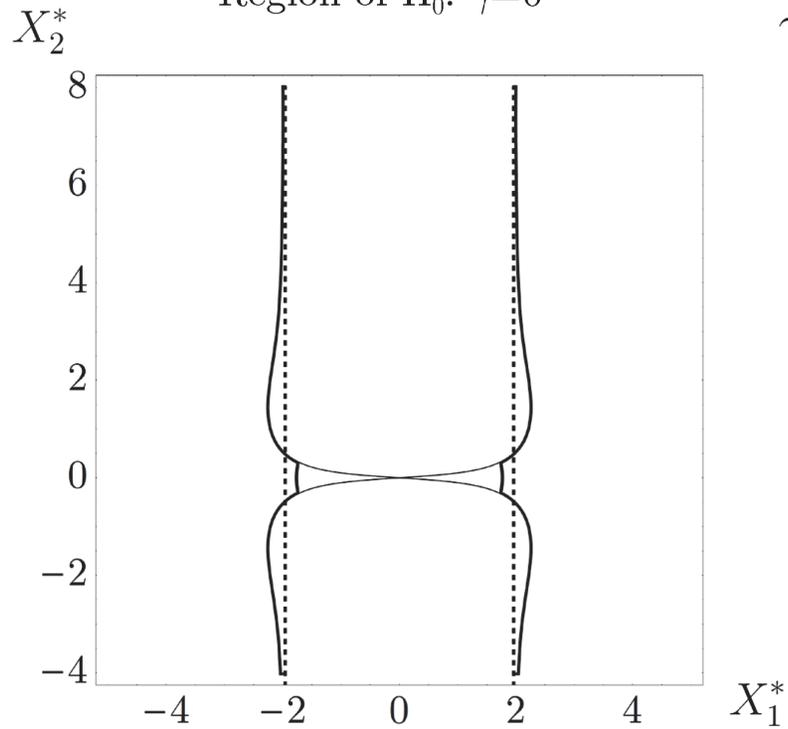
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## Weak Instruments

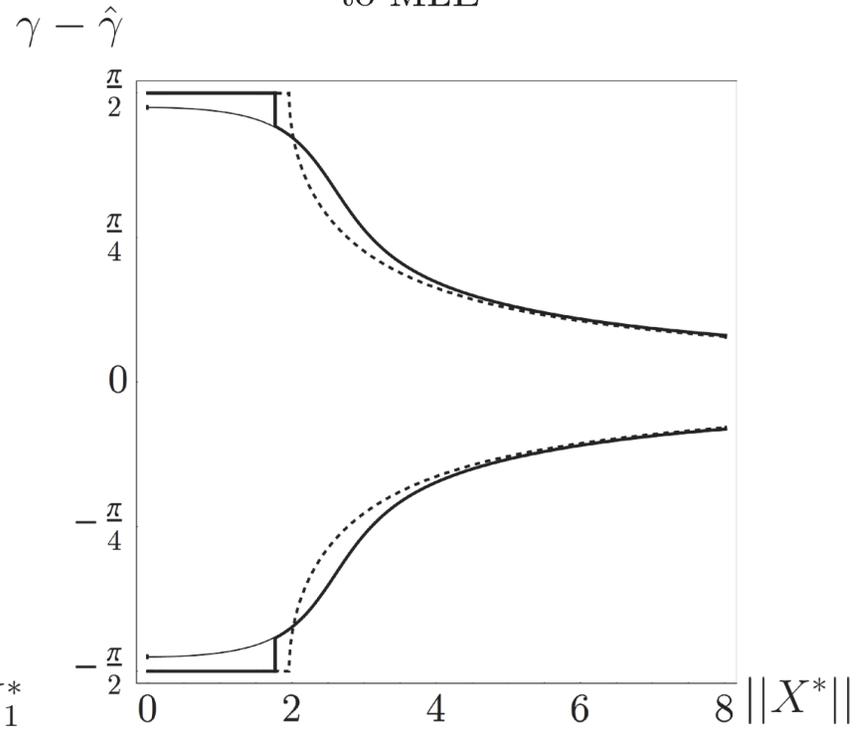
- Impose rotational invariance
- Specify prior  $\pi^0(\rho) \propto (100 + \rho)^{-1.1} \mathbf{1}[\rho \geq 0]$

# Weak Instruments

Boundaries of Rejection  
Region of  $H_0: \gamma=0$



Interval Bounds relative  
to MLE



Augmented Credible Set  
 AR CI  
 Credible Set

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## Concluding Remarks

- Alternative perspective on confidence sets in non-standard problems
- Also applicable to construction of predictive sets in forecasting problem under parameter uncertainty