

A NOTE ON THE DERIVATION OF
THEIL'S BLUS RESIDUALS

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In two papers the results of which are summarized in his Principles of Econometrics [1, Chapter 5], H. Theil has proposed a best linear unbiased (BLU) estimator of the residuals in a standard linear regression which has a scalar (S) covariance matrix (of the form $\sigma^2 I$). Theil's proof in the form of seven Lemma's [1, pp. 209-213] and other related theorems appears complicated and can be further motivated. This note provides a simple constructive proof of the BLUS residuals.

To set up the problem, consider n observations of a regression model with k explanatory variables

$$(1) \quad y = X\beta + \epsilon$$

where $E\epsilon\epsilon' = I\sigma^2$. For a linear estimator Cy to be unbiased, we have

$$(2) \quad E(Cy) = EC(X\beta + \epsilon) = CX\beta = 0 ,$$

which implies

$$(3) \quad CX = 0 .$$

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It follows from (3) [1, p. 207, Theorem 5.5] that the linear estimator can be written in three alternative forms

$$(4) \quad Cy = C\varepsilon = Ce$$

where e is the vector of residuals by the method of least squares

$$(5) \quad e = [I - X(X'X)^{-1}X']\varepsilon .$$

In order for $Cy = C\varepsilon$ to have a scalar covariance matrix

$$(6) \quad E(C\varepsilon\varepsilon'C') = CC'\sigma^2 = I\sigma^2 ,$$

we require

$$(7) \quad CC' = I .$$

The restriction (3), with X having k columns, is a set of k linear restrictions on the columns of C . Therefore, the rank of C cannot exceed $n-k$. We thus let C be an $(n-k) \times n$ matrix and Cy be an estimator of only $n-k$ components of ε . Accordingly Theil partitions model (1) as

$$(8) \quad \begin{bmatrix} y_0 \\ y_1 \end{bmatrix} = \begin{bmatrix} x_0 \\ x_1 \end{bmatrix} \beta + \begin{bmatrix} \varepsilon_0 \\ \varepsilon_1 \end{bmatrix} = \begin{bmatrix} x_0 \\ x_1 \end{bmatrix} b + \begin{bmatrix} e_0 \\ e_1 \end{bmatrix}$$

with X_0 being a nonsingular $k \times k$ matrix and X_1 being an $(n-k) \times k$ matrix. The linear estimator can be written as

$$(9) \quad Ce = C_0 e_0 + C_1 e_1,$$

C_1 being $(n-k) \times (n-k)$. Using $X_0' e_0 + X_1' e_1 = 0$, we have $e_0 = -(X_1 X_0^{-1})' e_1 = -Z' e_1$, where $Z = X_1 X_0^{-1}$. Further using $C_0 X_0 + C_1 X_1 = 0$, we have $C_0 = -C_1 X_1 X_0^{-1} = -C_1 Z$. (9) thus becomes

$$(10) \quad Ce = (-C_1 Z)(-Z' e_1) + C_1 e_1 = C_1 [I + ZZ'] e_1.$$

The constraint (7) for a scalar covariance matrix yields

$$(11) \quad C_0 C_0' + C_1 C_1' = C_1 Z Z' C_1' + C_1 C_1' = C_1 [I + ZZ'] C_1' = I.$$

The linear estimator is alternatively written as

$$(12) \quad C_0 \varepsilon_0 + C_1 \varepsilon_1 = -C_1 Z \varepsilon_0 + C_1 \varepsilon_1.$$

The vector of errors in using (12) to estimate ε_1 has a covariance matrix

$$(13) \quad \text{Cov}[-C_1 Z \varepsilon_0 + (C_1 - I) \varepsilon_1] = \sigma^2 [C_1 (I + ZZ') C_1' + I - C_1 - C_1'].$$

To find the matrix C_1 in the estimator (10) which minimizes the trace of (13), following Theil's definition of being "best," subject to the constraint (11) for a scalar covariance matrix, I would propose to form the Lagrangean expression

$$(14) \quad L = \text{tr}[C_1(I+ZZ')C_1' + I - C_1 - C_1'] - \text{tr} M[C_1(I+ZZ')C_1' - I]$$

where M is a symmetric $(n-k) \times (n-k)$ matrix of Lagrange multipliers, and differentiate with respect to C_1 . Using the differentiation rule $\partial \text{tr}(AB) / \partial \text{tr}(BA) / \partial A = B'$, we have

$$(15) \quad \frac{\partial L}{\partial C_1} = 2C_1(I+ZZ') - 2I - 2MC_1(I+ZZ') = 0$$

To solve equations (15) and (11) for the two unknowns C_1 and M , we post-multiply (15) by C_1' and use (11) to obtain

$$(16) \quad M = I - C_1'$$

Substituting (16) for M in (15), one gets

$$(17) \quad C_1' C_1 (I+ZZ') = I$$

Since C_1 is symmetric because M in (16) is symmetric, (17) implies $C_1^2 = (I+ZZ')^{-1}$ or

$$(18) \quad C_1 = (I+ZZ')^{-1/2}$$

which is the solution. Theil has written the solution in the form of $C_1 = PDP'$ where P is a matrix consisting of the characteristics vectors of $(I+ZZ')$ corresponding to the roots d_i^{-2} , D having diagonal elements d_i and $(I+ZZ') = PD^{-2}P'$. The reader may consult Theil [1] for further treatment of this topic.

REFERENCE

- [1] Theil, H: Principles of Econometrics. New York, John Wiley & Sons, Inc., 1971.